# STA 610L: MODULE 3.6

#### LOGISTIC MIXED EFFECTS MODEL (PART II)

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The dataset includes 2193 observations from one of eight surveys (the most recent CBS News survey right before the election) in the original full data.

Variable	Description				
org	cbsnyt = CBS/NYT				
bush	1 = preference for Bush Sr., 0 = otherwise				
state	1-51: 50 states including DC (number 9)				
edu	education: 1=No HS, 2=HS, 3=Some College, 4=College Grad				
age	1=18-29, 2=30-44, 3=45-64, 4=65+				
female	1=female, 0=male				
black	1=black, 0=otherwise				
region	1=NE, 2=S, 3=N, 4=W, 5=DC				
v_prev	average Republican vote share in the three previous elections (adjusted for home-state and home- region effects in the previous elections)				

Given that the data has a natural multilevel structure (through state and region), it makes sense to explore hierarchical models for this data.



Both voting turnout and preferences often depend on a complex combination of demographic factors.

In our example dataset, we have demographic factors such as biological sex, race, age, education, which we may all want to look at by state, resulting in  $2 \times 2 \times 4 \times 4 \times 51 = 3264$  potential categories of respondents.

We may even want to control for region, adding to the number of categories.

Clearly, without a very large survey (most political survey poll around 1000 people), we will need to make assumptions in order to even obtain estimates in each category.

We usually cannot include all interactions; we should therefore select those to explore (through EDA and background knowledge).

The data is in the file polls\_subset.txt on Sakai.



###### Load the data
polls\_subset <- read.table("data/polls\_subset.txt",header=TRUE)
str(polls\_subset)</pre>

head(polls\_subset)

## org	survey	bush	state	edu	age	female	black	region	v_prev
## 1 cbsnyt	9158	NA	7	3	1	1	0	1	0.5666333
## 2 cbsnyt	9158	1	39	4	2	1	0	1	0.5265667
## 3 cbsnyt	9158	0	31	2	4	1	0	1	0.5641667
## 4 cbsnyt	9158	0	7	3	1	1	0	1	0.5666333
## 5 cbsnyt	9158	1	33	2	2	1	0	1	0.5243666
## 6 cbsnyt	9158	1	33	4	4	1	Θ	1	0.5243666



summary(polls\_subset)

##	org	survey	bush	state
##			Min. :0.0000	
##	Class :characte	r 1st Qu.:9158	1st Qu.:0.0000	1st Qu.:14.00
##	Mode :characte	r Median :9158	Median :1.0000	Median :26.00
##		Mean :9158	Mean :0.5578	Mean :26.11
##		3rd Qu.:9158	3rd Qu.:1.0000	3rd Qu.:39.00
##		Max. :9158	Max. :1.0000	Max. :51.00
##			NA's :178	
##	edu	age	female	black
##	Min. :1.000	Min. :1.000	Min. :0.0000	Min. :0.00000
##	1st Qu.:2.000	1st Qu.:2.000	1st Qu.:0.0000	1st Qu.:0.00000
##	Median :2.000	Median :2.000	Median :1.0000	Median :0.00000
##	Mean :2.653	Mean :2.289	Mean :0.5887	Mean :0.07615
##	3rd Qu.:4.000	3rd Qu.:3.000	3rd Qu.:1.0000	3rd Qu.:0.00000
##	Max. :4.000	Max. :4.000	Max. :1.0000	Max. :1.00000
##				
##	region	v_prev		
##	Min. :1.000	Min. :0.1530		
##	1st Qu.:2.000	1st Qu.:0.5278		
##	Median :2.000	Median :0.5481		
##	Mean :2.431	Mean :0.5550		
##	3rd Qu.:3.000	3rd Qu.:0.5830		
##	Max. :5.000	Max. :0.6927		
##				



## [1] "AL" "AK" "AZ" "AR" "CA" "CO" "CT" "DE" "FL" "GA" "HI" "ID" "IL" "IN" "IA"
## [16] "KS" "KY" "LA" "ME" "MD" "MA" "MI" "MN" "MS" "MO" "MT" "NE" "NV" "NH" "NJ"
## [31] "NM" "NY" "NC" "ND" "OH" "OK" "OR" "PA" "RI" "SC" "SD" "TN" "TX" "UT" "VT"
## [46] "VA" "WA" "WV" "WI" "WY"

#In the polls data, DC is the 9th "state" in alphabetical order state\_abbr <- c (state.abb[1:8], "DC", state.abb[9:50]) polls\_subset\$state\_label <- factor(polls\_subset\$state,levels=1:51,labels=state\_abbr) rm(list = ls(pattern = "state")) #remove unnecessary values in the environment



###### View properties of the data
head(polls\_subset)

##		org	survey	bush	state	edu	age	female	black	region	v_prev	region_label
##	1	cbsnyt	9158	NA	7	3	1	1	0	1	56.66333	NE
##	2	cbsnyt	9158	1	39	4	2	1	0	1	52.65667	NE
##	3	cbsnyt	9158	Θ	31	2	4	1	Θ	1	56.41667	NE
##	4	cbsnyt	9158	0	7	3	1	1	0	1	56.66333	NE
##	5	cbsnyt	9158	1	33	2	2	1	0	1	52.43666	NE
##	6	cbsnyt	9158	1	33	4	4	1	Θ	1	52.43666	NE
<pre>## edu_label age_label state_label</pre>												
##	1	Some Co	ollege	18	3-29		C	Т				
##	2	College	e Grad	30	9-44		F	PA				
##	3		HS		65+		Ν	13				
##	4	Some Co	ollege	18	3-29		C	Т				
##	5		HS	30	9-44		Ν	IY				
##	6	College	e Grad		65+		Ν	IY				

dim(polls\_subset)

## [1] 2193 14



###### View properties of the data
str(polls\_subset)

##	'data.frame': 2193	obs. of 14 variables:
##	\$ org : chr	"cbsnyt" "cbsnyt" "cbsnyt"
##	\$ survey : int	9158 9158 9158 9158 9158 9158 9158 9158
##	\$ bush : int	NA 1 0 0 1 1 1 1 0 0
##	\$ state : int	7 39 31 7 33 33 39 20 33 40
##	\$ edu : int	3 4 2 3 2 4 2 2 4 1
##	\$ age : int	1 2 4 1 2 4 2 4 3 3
##	\$ female : int	1 1 1 1 1 0 1 0 0
##	\$ black : int	$0 0 0 0 0 0 0 0 0 \dots$
##	<pre>\$ region : int</pre>	1 1 1 1 1 1 1 1 1
##	\$ v_prev : num	56.7 52.7 56.4 56.7 52.4
##	<pre>\$ region_label: Fact</pre>	or w/ 5 levels "NE","S","N","W",: 1 1 1 1 1 1 1 1 1
##	<pre>\$ edu_label : Fact</pre>	cor w/ 4 levels "No HS","HS","Some College",: 3 4 2 3 2 4 2 2 4 1
##	<pre>\$ age_label : Fact</pre>	cor w/ 4 levels "18-29","30-44",: 1 2 4 1 2 4 2 4 3 3
##	<pre>\$ state_label : Fact</pre>	or w/ 51 levels "AL","AK","AZ",: 7 39 31 7 33 33 39 20 33 40



I will not do any meaningful EDA here.

I expect you to be able to do this yourself.

Let's just take a look at the amount of data we have for "bush" and the age:edu interaction.

###### Exploratory data analysis
table(polls\_subset\$bush) #well split by the two values

## ## 0 1 ## 891 1124

table(polls\_subset\$edu,polls\_subset\$age)

##
## 1 2 3 4
## 1 44 42 67 96
## 2 232 283 223 116
## 3 141 205 99 54
## 4 119 285 125 62



As a start, we will consider a simple model with fixed effects of race and sex and a random effect for state (50 states + the District of Columbia).

$$egin{aligned} ext{bush}_{ij} | oldsymbol{x}_{ij} &\sim ext{Bernoulli}(\pi_{ij}); \quad i=1,\ldots,n; \quad j=1,\ldots,J=51; \ &\log\left(rac{\pi_{ij}}{1-\pi_{ij}}
ight) = eta_0 + b_{0j} + eta_1 ext{female}_{ij} + eta_2 ext{black}_{ij}; \ &b_{0j} \sim N(0,\sigma^2). \end{aligned}$$

In R, we have



```
## Generalized linear mixed model fit by maximum likelihood (Laplace
    Approximation) [glmerMod]
##
## Family: binomial ( logit )
## Formula: bush ~ black + female + (1 | state label)
     Data: polls subset
##
##
                BIC logLik deviance df.resid
##
       AIC
##
    2666.7
             2689.1 -1329.3 2658.7
                                         2011
##
## Scaled residuals:
##
      Min
              10 Median
                              30
                                     Max
## -1.7276 -1.0871 0.6673 0.8422 2.5271
##
## Random effects:
## Groups
                         Variance Std.Dev.
               Name
## state label (Intercept) 0.1692 0.4113
## Number of obs: 2015, groups: state label, 49
##
## Fixed effects:
##
              Estimate Std. Error z value Pr(|z|)
## (Intercept) 0.44523 0.10139 4.391 1.13e-05
## black -1.74161 0.20954 -8.312 < 2e-16
         -0.09705 0.09511 -1.020
## female
                                         0.308
##
## Correlation of Fixed Effects:
##
         (Intr) black
## black -0.119
## female -0.551 -0.005
```



Looks like we dropped some NAs.

c(sum(complete.cases(polls\_subset)),sum(!complete.cases(polls\_subset)))

## [1] 2015 178

Not ideal; we'll learn about methods for dealing with missing data soon.

Interpretation of results:

- For a fixed state (or across all states), a non-black male respondent has odds of  $e^{0.45} = 1.57$  of supporting Bush.
- For a fixed state and sex, a black respondent as e<sup>-1.74</sup> = 0.18 times (an 82% decrease) the odds of supporting Bush as a non-black respondent; you are much less likely to support Bush if your race is black compared to being non-black.
- For a given state and race, a female respondent has  $e^{-0.10} = 0.91$  (a 9% decrease) times the odds of supporting Bush as a male respondent. However, this effect is not actually statistically significant!



The state-level standard deviation is estimated at 0.41, so that the states do vary some, but not so much.

I expect that you will be able to interpret the corresponding confidence intervals.

## Computing profile confidence intervals ...

##		2.5 %	97.5 %
##	.sig01	0.2608567	0.60403428
##	(Intercept)	0.2452467	0.64871247
##	black	-2.1666001	-1.34322366
##	female	-0.2837100	0.08919986



We can definitely fit a more sophisticated model that includes other relevant survey factors, such as

- region
- prior vote history (note that this is a state-level predictor),
- age, education, and the interaction between them.

Given the structure of the data, it makes sense to include region as a second grouping variable.

We will return to this soon.



For now, let's just fit two models, one with the main effects for age and education, and the second with the interaction between them.

```
## Generalized linear mixed model fit by maximum likelihood (Laplace
##
    Approximation) [glmerMod]
   Family: binomial (logit)
##
## Formula: bush ~ black + female + edu label + age label + (1 | state label)
##
     Data: polls subset
##
##
       AIC
                BIC
                    logLik deviance df.resid
##
    2662.2 2718.3 -1321.1 2642.2
                                         2005
##
## Scaled residuals:
##
      Min
               10 Median
                              30
                                     Max
## -1.8921 -1.0606 0.6420 0.8368 2.7906
##
## Random effects:
##
   Groups
               Name
                          Variance Std.Dev.
## state label (Intercept) 0.1738 0.4168
## Number of obs: 2015, groups: state label, 49
##
## Fixed effects:
##
                        Estimate Std. Error z value Pr(|z|)
## (Intercept)
                      0.31206
                                   0.19438 1.605 0.10841
## black
                       -1.74378
                                   0.21124 -8.255 < 2e-16
## female
                       -0.09681
                                   0.09593 -1.009 0.31289
## edu labelHS
                        0.23282
                                   0.16569 1.405 0.15998
## edu_labelSome College 0.51598
                                   0.17921 2.879 0.00399
## edu labelCollege Grad 0.31585
                                   0.17454 1.810 0.07036
## age label30-44
                    -0.29222
                                   0.12352 -2.366 0.01800
                   -0.06744
-0.22509
## age_label45-64
                                   0.13738 -0.491 0.62352
## age_label65+
                       -0.22509
                                   0.16142 -1.394 0.16318
```

#### Can you interpret the results?



```
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, :
## Model failed to converge with max|grad| = 0.00802313 (tol = 0.002, component 1)
```

Looks like we have a convergence issue. These convergence issues are really common when using glmer.

Here we have so many parameters to estimate from the interaction terms edu\_label\*age\_label (16 actually), and it looks like that's causing a problem.

Now, there are a few potential reasons and fixes for this problem (see this link) but we'll see how we can actually take advantage of the properties of our hierarchical model to get around the issue.

Side note: if you suspect your design matrix is not full rank, you can do a quick check using the rankMatrix function in the Matrix package.



#### QUICK NOTE ON ESTIMATION

ML estimation is carried out typically using adaptive Gaussian quadrature.

To improve accuracy, many packages (default is usually Laplace approximation) increase the number of quadrature points to be greater than one.

Note that some software packages (including the glmer function in the lme4 package) require Laplace approximation with Gaussian quadrature if the number of random effects is more than 1 for the sake of computational efficiency.

The main point though is that it is possible to tweak the approximation, and specifically the optimizer, in the glmer function, so that the usual go-to solution for getting around convergence issues is to simply change the optimizer.

Read more about the BOBYQA optimizer in particular at your leisure.

**My take:** as I have mentioned before, hierarchical modeling is one of the areas where leaning Bayesian is a huge plus; not having to deal with convergence issues is one of them.

First, let's go back to the model without the interaction but then try to control for

- region (since states are nested within regions)
- prior vote history (our state-level predictor),

#### We have

```
## Warning in checkConv(attr(opt, "derivs"), opt$par, ctrl = control$checkConv, :
## Model failed to converge with max|grad| = 0.0437183 (tol = 0.002, component 1)
```

which also does not converge.



We are unable to include education and age in this version of the model. Could be that we have too little  $bush_i = 1$  or 0 values for certain combinations? You should check!

As mentioned before, we can actually take advantage of the properties of our hierarchical model to get around the issue.

How about we treat those as varying/random effects instead? Let's try

This runs fine. Here we are able to borrow information for the combinations of those variables with insufficient data, and that helps a ton!

This is more of an adhoc fix, but it often works really well in practice.

**Side note:** ideally, we should be much more careful with building the model (for example, do we really need to include region?).



summary(model3)

```
## Generalized linear mixed model fit by maximum likelihood (Laplace
    Approximation) [glmerMod]
##
##
   Family: binomial (logit)
## Formula:
## bush ~ black + female + v prev + (1 | state label) + (1 | region label) +
      (1 | edu label:age label)
##
     Data: polls subset
##
##
                    logLik deviance df.resid
##
       AIC
                BIC
    2644.0
            2683.3 -1315.0 2630.0
##
                                          2008
##
## Scaled residuals:
##
      Min
               10 Median
                               30
                                      Max
## -1.8404 -1.0430 0.6478 0.8405 2.7528
##
## Random effects:
## Groups
                       Name
                                   Variance Std.Dev.
## state label
                       (Intercept) 0.03768 0.1941
## edu label:age label (Intercept) 0.02993 0.1730
## region label
                       (Intercept) 0.02792 0.1671
## Number of obs: 2015, groups:
## state_label, 49; edu_label:age_label, 16; region_label, 5
##
## Fixed effects:
##
              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -3.50658
                         1.03365 -3.392 0.000693
## black
              -1.74530
                          0.21090
                                  -8.275 < 2e-16
## female
              -0.09956
                          0.09558
                                  -1.042 0.297575
               0.07076
## v prev
                          0.01853
                                   3.820 0.000134
##
## Correlation of Fixed Effects:
##
         (Intr) black female
## black -0.036
## female -0.049 -0.004
## v_prev -0.992 0.027 -0.006
```



Remember that in the first model, the state-level standard deviation was estimated as 0.41. Looks like we are now able to separate that (for the most part) into state and region effects.

Interpretation of results:

- For a fixed state, education and age bracket, a non-black male respondent with zero prior average Republican vote share, has odds of  $e^{-3.51} = 0.03$  of supporting Bush (no one really has 0 value for v\_prev).
- For a fixed state, sex, education level, age bracket and zero prior average Republican vote share, a black respondent has  $e^{-1.75} = 0.17$  times (an 83% decrease) the odds of supporting Bush as a non-black respondent, which is about the same as before.
- For each percentage point increase in prior average Republican vote share, residents of a given state, race, sex, education level age bracket have  $e^{0.07} = 1.07$  times the odds of supporting Bush.



# WHAT'S NEXT?

Move on to the readings for the next module!

